

Multivariate linear regression

Setup

$$y_i = x_i^T \beta^* + \epsilon_i, i = 1, \dots, n$$

where

- x_i is the vector of explanatory variables or covariates
- y_i is the response/dependent variable
- $\beta^* = (a^*, b^T)^T$ (a^* is the intercept)
- $\epsilon_{i=1, \dots, n}$: noise terms

Then, the least squares estimator (LSE) of $\hat{\beta}$ is the minimizer of the sum of errors squared:

$$\hat{\beta} = \underset{\beta \in \mathbb{R}^p}{\operatorname{argmin}} \sum_{i=1}^n (y_i - x_i^T \beta)^2$$

Matrix form

- $y = (y_1, \dots, y_n)^T \in \mathbb{R}_n$ (vector of observations)
- X be the design matrix whose rows are $x_1^T, \dots, x_n^T$ (aka the design matrix, $n \times p$)
- $\epsilon = (\epsilon_1, \dots, \epsilon_n)^T \in \mathbb{R}^n$ (vector of noise)
- Then, $y = X\beta^* + \epsilon$. β^* is the unknown model parameter.

The least squares estimator of the unknown model parameter $\hat{\beta}$ is:

$$\hat{\beta} = \underset{\beta \in \mathbb{R}^p}{\operatorname{argmin}} \|y - X\beta\|_2^2$$

Each row of X represents one set of explanatory variables, and the corresponding row/element of y represents the response variable for that set of explanatory variables. The corresponding row/element of ϵ represents the error between the true response variable and the value predicted by the regression model.

X is an $n \times p$ matrix, where n is the number of observations, and p is the number of covariates, including one constant covariate.

Evaluating the least-squares estimator

By setting the gradient of the sum of errors squared to zero, we find that the LSE $\hat{\boldsymbol{\beta}}$ must satisfy:

$$\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} = \mathbf{X}^T \mathbf{y}$$

To isolate $\hat{\boldsymbol{\beta}}$, we can multiply both sides by $(\mathbf{X}^T \mathbf{X})^{-1}$ from the left. To do this, $\mathbf{X}^T \mathbf{X}$ must be invertible. $\mathbf{X}$ having rank equal to the number of covariates will guarantee that $\mathbf{X}^T \mathbf{X}$ is invertible.

If $\text{rank}(\mathbf{X}) < p$, where p is the number of covariates, there will be an infinite collection of estimators that satisfy the least-squares condition.

If $\text{rank}(\mathbf{X}) = p$, there will be a unique LSE $\hat{\boldsymbol{\beta}}$.

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

Deterministic design

When we use deterministic design, we make the following assumptions:

- $\mathbf{X}$ is deterministic, and $\text{rank}(\mathbf{X}) = p$
- $\epsilon_1, \dots, \epsilon_n$ are i.i.d. (The model is homoscedastic)
- The noise vector $\boldsymbol{\epsilon}$ is Gaussian. $\boldsymbol{\epsilon} \sim \mathcal{N}(0, \sigma^2 \mathbf{I}_n)$

$$\mathbf{y} = \mathbf{X} \boldsymbol{\beta} + \boldsymbol{\epsilon}$$

Implications

- This way, the only random element in the equation for the response variable $\mathbf{y}$ is the noise $\boldsymbol{\epsilon}$.
- The response variable $\mathbf{y}$ is therefore a Gaussian random variable.
- The LSE $\hat{\boldsymbol{\beta}}$ is also a Gaussian random variable.

LSE properties

The LSE is equal to the maximum likelihood estimator (MLE).

Distribution

$\hat{\beta} \sim \mathcal{N}_p(\beta^*, \sigma^2(\mathbb{X}^T \mathbb{X})^{-1})$ The distribution of the LSE $\hat{\beta}$ is a p -dimensional Gaussian with mean β^* and variance $\sigma^2(\mathbb{X}^T \mathbb{X})^{-1}$.

Quadratic risk

$$\mathbb{E}[\|\hat{\beta} - \beta\|_2^2] = \text{tr}(\sigma^2(\mathbb{X}^T \mathbb{X})^{-1})$$

The quadratic risk is defined as the typical error in the LSE $\hat{\beta}$ compared to the true parameter β .

$\text{tr}(\mathbb{X})$ is the trace, defined as the sum of elements on the main diagonal of $\mathbb{X}$.

Prediction error

$$\mathbb{E}[\|\text{Y} - \mathbb{X} \hat{\beta}\|_2^2] = \sigma^2(n-p)$$

The prediction error is defined as the typical error between model predictions $\mathbb{X} \hat{\beta}$ and observations Y .

Variance estimator

Unbiased estimator of σ^2 : $\hat{\sigma}^2 = \frac{\|\text{Y} - \mathbb{X} \hat{\beta}\|_2^2}{n-p} = \frac{1}{n-p} \sum_{i=1}^n \hat{\epsilon}_i^2$

Theorem

$$(n-p) \frac{\hat{\sigma}^2}{\sigma^2} \sim \chi_{n-p}^2 \quad \hat{\beta} \perp \hat{\sigma}^2$$

Significance testing

Hypothesis testing setup example:

$$H_0: \beta_j = 0 \quad H_1: \beta_j \neq 0$$

If γ_j is the j th diagonal coefficient of $(\mathbb{X}^T \mathbb{X})^{-1}$

$$\frac{\hat{\beta}_j - \beta_j}{\sqrt{\hat{\sigma}^2 \gamma_j}} \sim t_{n-p}$$

The test statistic is $T_n(j) = \frac{\hat{\beta}_j}{\sqrt{\hat{\sigma}^2 \gamma_j}}$.

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